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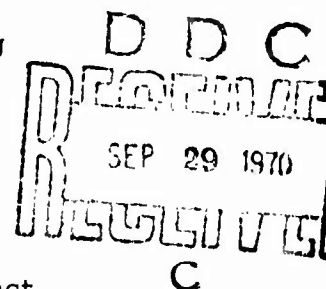
THEMIS SIGNAL ANALYSIS STATISTICS RESEARCH PROGRAM

CORRELATION BETWEEN TWO HOTELLING'S  $T^2$

by

A. M. Kshirsagar\* and John C. Young

Technical Report No. 79  
Department of Statistics THEMIS Contract



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**Department of Statistics**  
**Southern Methodist University**  
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# CORRELATION BETWEEN TWO HOTELLING'S $T^2$

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## 1. Introduction:

Let  $B$  be a  $p \times p$  symmetric matrix having the Wishart distribution

$$(1.1) \quad W_p(B|I|f)dB = C_{pf}|B|^{(f-p-1)/2} e^{-1/2 \text{tr} B} dB,$$

where

$$(1.2) \quad C_{pf}^{-1} = 2^{fp/2} \pi^{p(p-1)/4} \prod_{i=1}^p \Gamma\left(\frac{f+1-i}{2}\right),$$

and  $dB$  stands for the product of the differentials of the  $p(p+1)/2$  distinct elements of  $B$ . Let  $\underline{x}$  and  $\underline{y}$  be two vector variables of  $p$  components, distributed independently of  $B$ , and also independently of each other, as

$$(1.3) \quad \frac{1}{(2\pi)^{p/2}} e^{-1/2 \underline{x}'\underline{x}} d\underline{x},$$

and

$$(1.4) \quad \frac{1}{(2\pi)^{p/2}} e^{-1/2 \underline{y}'\underline{y}} d\underline{y}$$

respectively. While considering the problem of multivariate statistical outliers, Wilks (1963) used statistics of the type,

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$$(1.5) \quad r = |B + \underline{y}\underline{y}'|/|B + \underline{x}\underline{x}' + \underline{y}\underline{y}'|, \quad s = |B + \underline{x}\underline{x}'|/|B + \underline{x}\underline{x}' + \underline{y}\underline{y}'|.$$

He has remarked that the exact distribution (joint) of  $r$  and  $s$  is complicated and has given the expected values, variances and covariance of  $r$  and  $s$ . Unfortunately, his expressions for the variance and covariance are in error. The purpose of this note is to derive the exact joint distribution of  $r$  and  $s$  and to give correct expressions for the moments.

## 2. Joint distribution:

In the joint distribution of  $B$ ,  $\underline{x}$  and  $\underline{y}$ , make the following transformation

$$(2.1) \quad \begin{aligned} A &= B + \underline{x}\underline{x}' + \underline{y}\underline{y}', \\ \underline{u} &= A^{-1/2} \underline{x}, \\ \underline{v} &= A^{-1/2} \underline{y}, \end{aligned}$$

where  $A^{-1/2}$  is any matrix such that  $A^{-1/2} \cdot A^{-1/2} = A^{-1}$ . The Jacobian of transformation from  $B$  to  $A$  is 1 and that from  $\underline{x}$  to  $\underline{u}$  or  $\underline{y}$  to  $\underline{v}$  is  $|A|^{1/2}$  and hence, the joint distribution of  $A$ ,  $\underline{u}$  and  $\underline{v}$  comes out as

$$(2.2) \quad \frac{C_{pf}}{(2\pi)^p} |A|^{\frac{(f+2)-p-1}{2}} e^{-1/2 \operatorname{tr} A} \cdot |I - \underline{u}\underline{u}' - \underline{v}\underline{v}'|^{\frac{f-p-1}{2}} dA d\underline{u} d\underline{v},$$

$$\text{as } |B| = |A - A^{1/2} \underline{u}\underline{u}' A^{1/2} - A^{1/2} \underline{v}\underline{v}' A^{1/2}| = |A| |I - \underline{u}\underline{u}' - \underline{v}\underline{v}'|.$$

This shows that  $A$  has a Wishart distribution of  $f+2$  degrees of freedom and is independent of  $\underline{u}$  and  $\underline{v}$ . Splitting the constant suitably, the joint distribution of  $\underline{u}$  and  $\underline{v}$  is

$$(2.3) \quad \frac{\Gamma(f+1)}{(2\pi)^p \Gamma(f-p+1)} |I - \underline{u}\underline{u}' - \underline{v}\underline{v}'|^{\frac{f-p-1}{2}} d\underline{u} d\underline{v}.$$

Observe that the statistics  $r$ ,  $s$  of Wilks are given by

$$(2.4) \quad r = \frac{|B + \underline{y}\underline{y}'|}{|B + \underline{x}\underline{x}' + \underline{y}\underline{y}'|} = \frac{|A - \underline{x}\underline{x}'|}{|A|} = |I - \underline{u}\underline{u}'| = 1 - \underline{u}'\underline{u},$$

and

$$(2.5) \quad s = \frac{|B+xx'|}{|B+xx'+yy'|} = 1 - \underline{v}'\underline{v} \quad .$$

Also observe that in (2.3)

$$(2.6) \quad \begin{aligned} |I - \underline{u}\underline{u}' - \underline{v}\underline{v}'| &= (1 - \underline{u}'\underline{u})(1 - \underline{v}'\underline{v}) - (\underline{u}'\underline{v})^2 \\ &= rs - (\underline{u}'\underline{v})^2 \quad . \end{aligned}$$

In (2.3), transform from  $\underline{v}$  to  $\underline{w} = [w_1, w_2, \dots, w_p]'$ , by an orthogonal transformation

$$(2.7) \quad \underline{w} = L\underline{v} \quad ,$$

where

$L$  is a  $p \times p$  orthogonal matrix, whose last row is  $\underline{u}'/\sqrt{\underline{u}'\underline{u}}$ . The Jacobian of this transformation is  $|L| = 1$  and  $\underline{v}'\underline{v} = \underline{w}'\underline{w} = 1-s$ . Also

$$(2.8) \quad \underline{u}'\underline{v} = \underline{u}'L'L\underline{v} = [0 \dots 0, \sqrt{\underline{u}'\underline{u}}] \underline{w} = \sqrt{1-r} \cdot w_p \quad .$$

The joint distribution of  $\underline{u}$  and  $\underline{w}$  is, therefore,

$$(2.9) \quad \frac{\Gamma(f+1)}{(2\pi)^p \Gamma(f-p+1)} \left\{ rs - (1-r)w_p^2 \right\}^{\frac{f-p-1}{2}} d\underline{u}d\underline{w} \quad .$$

From  $\underline{u}$ , transform to  $r = 1 - \underline{u}'\underline{u}$  and  $p-1$  other variables

$\phi_1, \phi_2, \dots, \phi_{p-1}$  by

$$(2.10) \quad \begin{aligned} u_1 &= (1-r)^{1/2} \cos\phi_1 \cos\phi_2 \dots \cos\phi_{p-1} \quad , \\ u_j &= (1-r)^{1/2} \cos\phi_1 \cos\phi_2 \dots \cos\phi_{p-j} \sin\phi_{p-j+1} \\ &\quad (j=2, 3, \dots, p) \end{aligned}$$

Similarly, transform from  $\underline{w}$  to  $s = 1 - \underline{w}'\underline{w}$  and  $p-1$  other variables

$\theta_1, \theta_2, \dots, \theta_{p-1}$  by

$$(2.11) \quad \begin{aligned} v_1 &= (1-s)^{1/2} \cos\theta_1 \cos\theta_2 \dots \cos\theta_{p-1} \quad , \\ v_j &= (1-s)^{1/2} \cos\theta_1 \cos\theta_2 \dots \cos\theta_{p-j} \sin\theta_{p-j+1} \quad . \\ &\quad (j=2, 3, \dots, p) \end{aligned}$$

The Jacobian of transformation from  $\underline{u}$  to  $r, \phi_1, \dots, \phi_{p-1}$  is

$$\frac{1}{2}(1-r)^{\frac{1}{2}p-1} \prod_{i=1}^{p-2} \cos \phi_i^{p-i-1}$$

and a similar expression in  $s$  and  $\theta_1$  for the Jacobian of transformation from  $\underline{w}$  to  $s$  and the  $\theta$ 's. Now  $\theta_{p-1}$  and  $\phi_{p-1}$  vary from 0 to  $2\pi$ , the other  $\theta$ 's and  $\phi$ 's vary from  $-\pi/2$  to  $\pi/2$  while  $r$  and  $s$  vary from 0 to 1. Integrating out all the  $\phi$ 's and all  $\theta$ 's except  $\theta_1$ , we obtain the joint distribution of  $r, s$  and  $\theta_1$  as

$$(2.12) \quad \frac{\Gamma(f+1)}{4\pi\Gamma(p-1)\Gamma(f-p+1)} \{rs - (1-r)(1-s)\sin^2\theta\}^{\frac{f-p-1}{2}} \cos^{p-2}\theta \, dr ds d\theta$$

where  $\theta_1$  is replaced by  $\theta$ .

The joint distribution of  $r, s$  alone can now be obtained by integrating out  $\theta$  but this does not seem to yield a manageable expression, as the bracket in (2.12) will have to be expanded in a series.

### 3. Moments of $r, s$ .

Only the product moment of  $r$  and  $s$  is difficult to obtain. The mean and variance of  $r$  (or  $s$ ) can be very easily obtained from the marginal distribution of  $r$ , which is related to the well-known Hotelling's

$T^2$  by  $r = \frac{1}{1 + \left(\frac{T^2}{f+1}\right)}$ . In the joint distribution of  $\underline{u}$  and  $\underline{v}$ , given by

(2.3), if we transform to  $\underline{z} = [z_1, \dots, z_p]'$  from  $\underline{v}$  by

$$(3.1) \quad \underline{v} = (I - \underline{u}\underline{u}')^{1/2} \underline{z},$$

we shall find that  $\underline{u}$  and  $\underline{z}$  are independently distributed as

$$(3.2) \quad K(\underline{u}|f) d\underline{u} = \frac{f}{\pi^{p/2}(f-p)} \cdot \frac{\Gamma(f/2)}{\frac{\pi}{2}(f-p)} |I - \underline{u}\underline{u}'|^{\frac{f-p}{2}} d\underline{u}$$

$$(3.3) \quad \text{and } K(\underline{z}|f-1) d\underline{z}, \text{ respectively.}$$

From (3.2), one can easily show that

$$(3.4) \quad E(r^h) = E(1 - \underline{u}'\underline{u})^h = E|I - \underline{u}\underline{u}'|^h \\ = \frac{f(f+2h-p)}{(f-p)(f+2h)} \cdot \frac{\Gamma(\frac{f-p}{2} + h)\Gamma(\frac{f}{2})}{\Gamma(\frac{f}{2} + h)\Gamma(\frac{f-p}{2})}$$

This will also be the  $h^{\text{th}}$  moment of  $s$  by symmetry. This leads to

$$(3.5) \quad E(r) = \frac{f-p+2}{f+2}, \quad v(r) = \frac{2p(f-p+2)}{(f+2)^2(f+4)}.$$

$$\text{Now } \text{Cov}(r, s) = E\{(1 - \underline{u}'\underline{u})(1 - \underline{v}'\underline{v})\} - E(r)E(s)$$

$$= E\{(1 - \underline{u}'\underline{u})[1 - \underline{z}(I - \underline{u}\underline{u}')\underline{z}]\} - \{E(r)\}^2 \quad \text{by (3.1)}$$

$$= E(r) - E\{r[\underline{z}'\underline{z} - (\underline{z}'\underline{u})^2]\} - \{E(r)\}^2$$

$$(3.6) \quad = E(r) - E(r)E(\underline{z}'\underline{z}) + E\{r(\underline{z}'\underline{u})^2\} - \{E(r)\}^2,$$

as  $\underline{z}$  and  $r$  are independent. Since  $\underline{z}$  has the same distribution as  $\underline{u}$  with  $f$  changed  $f-1$ ,

$$(3.7) \quad E(\underline{z}'\underline{z}) = 1 - E(1 - \underline{u}'\underline{u}) \text{ with } f \text{ replaced by } f-1 \\ = \frac{p}{f+1}$$

Hence (3.6) reduces to

$$(3.8) \quad \text{Cov}(r, s) = \frac{-p(f-p+2)}{(f+1)(f+2)^2} + E\{r(\underline{z}'\underline{u})^2\}.$$

Now

$$(3.9) \quad E\{r(\underline{z}'\underline{u})^2\} = \int (1 - \underline{u}'\underline{u}) (\underline{z}'\underline{u})^2 K(\underline{u}|f) K(\underline{z}|f-1) d\underline{u} d\underline{z}$$

where the integration is over the range of values of  $\underline{u}$  and  $\underline{z}$  such that  $\underline{u}'\underline{u} \leq 1$ ,  $\underline{z}'\underline{z} \leq 1$ . Transform from  $\underline{z}$  to  $\underline{\xi} = [\xi_1, \dots, \xi_p]$  by the transformation

$$\underline{\xi} = L\underline{z},$$

where  $L$  is already defined to be a  $p \times p$  orthogonal matrix, whose last row is  $\underline{u}'/\sqrt{\underline{u}'\underline{u}}$ . Then,

$$\underline{z}'\underline{u} = \underline{z}'L'L\underline{u} = \underline{\xi}'L\underline{u} = \xi_p \sqrt{\underline{u}'\underline{u}} = (1-r)^{1/2} \xi_p.$$

Hence (3.9) reduces to

(3.10)  $\int r(1-r)K(\underline{u}|f)d\underline{u} = \int_0^1 K(\underline{t}|f-1)d\underline{t} = K(r-r^2) \cdot \frac{1}{p} K(\underline{t}'\underline{t})$ , due to symmetry of the distribution of  $\underline{t}$ . Now  $\underline{t}$  has the same distribution as  $\underline{u}$  with  $f$  replaced by  $f-1$  and hence finally, (3.10) reduces to

$$\frac{p(f-p+2)}{(f+4)(f+2)} \cdot \frac{1}{f+1}.$$

The covariance between  $r$  and  $s$ , therefore, is (from (3.5))

$$(3.11) \quad \frac{-2p(f-p+2)}{(f+1)(f+2)^2(f+4)}.$$

#### Remarks:

Wilks considers a sample of size  $n$  and has a Wishart matrix based on  $n-1$  degrees of freedom as deviations are from the sample means. He then removes two observations as outliers and thus his  $(n-1)-2$  corresponds to our  $f$ . His  $E(r)$  agrees with our result, with this correspondence but the other moments are in error.

#### Reference

Wilks, S. S. (1963). "Multivariate Statistical Outliers," Sankhyā, Vol. 25, p. 407-426.



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8. SUPPLEMENTARY NOTES		9. SPONSORING MILITARY ACTIVITY Office of Naval Research	
10. ABSTRACT If $B$ is a Wishart matrix and $\underline{x}$ , $\underline{y}$ are two vectors of $p$ components each having a multinormal distribution and if all these quantities are independently distributed, the joint distribution of the two statistics $r = \frac{ B + \underline{y}\underline{y}' }{ B + \underline{x}\underline{x}' + \underline{y}\underline{y}' } \quad \text{and} \quad s = \frac{ B + \underline{x}\underline{x}' }{ B + \underline{x}\underline{x}' + \underline{y}\underline{y}' }$ is derived in this paper. The correlation between $r$ and $s$ is also obtained. $r$ and $s$ are related to Hotelling's $T^2$ and are useful in problems of testing multivariate outliers.			